

$$HOUS = \beta_1 + \beta_2 DPI + \beta_3 PRELHOUS + u$$

In this sequence we will make an initial exploration of the determinants of aggregate consumer expenditure on housing services using the Demand Functions data set.

$$HOUS = \beta_1 + \beta_2 DPI + \beta_3 PRELHOUS + u$$

***HOUS*** is aggregate consumer expenditure on housing services and ***DPI*** is aggregate disposable personal income. Both are measured in \$ billion at 2000 constant prices.

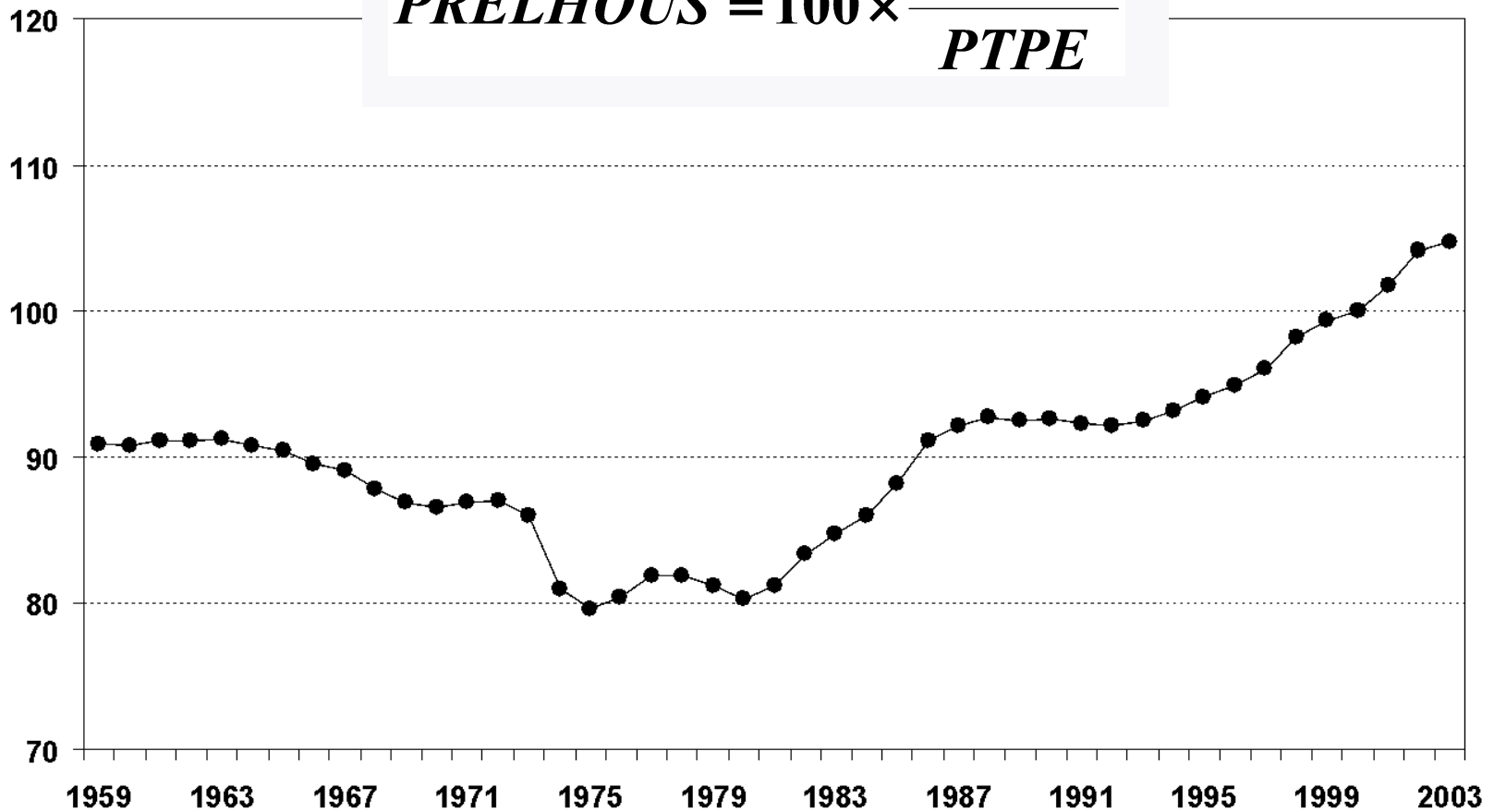
$$HOUS = \beta_1 + \beta_2 DPI + \beta_3 PRELHOUS + u$$

$$PRELHOUS = 100 \times \frac{PHOUS}{PTPE}$$

***PRELHOUS*** is a relative price index for housing services constructed by dividing the nominal price index for housing services by the price index for total personal expenditure.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

$$PRELHOUS = 100 \times \frac{PHOUS}{PTPE}$$



Here is a plot of  $PRELHOUS$  for the sample period, 1959–2003.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

=====  
Dependent Variable: HOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45  
=====

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	334.6657	37.26625	8.980396	0.0000
DPI	0.150925	0.001665	90.65785	0.0000
PRELHOUS	-3.834387	0.460490	-8.326764	0.0000

=====

R-squared	0.996722	Mean dependent var	630.2830
Adjusted R-squared	0.996566	S.D. dependent var	249.2620
S.E. of regression	14.60740	Akaike info criteri	8.265274
Sum squared resid	8961.801	Schwarz criterion	8.385719
Log likelihood	-182.9687	F-statistic	6385.025
Durbin-Watson stat	0.337638	Prob(F-statistic)	0.000000

=====

Here is the regression output using EViews. It was obtained by loading the workfile, clicking on Quick, then on Estimate, and then typing HOUS C DPI PRELHOUS in the box. Note that in EViews you must include C in the command if your model has an intercept.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: HOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	334.6657	37.26625	8.980396	0.0000
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Durbin-Watson stat	0.337638	Prob(F-statistic)	0.000000

We will start by interpreting the coefficients. The coefficient of *DPI* indicates that if aggregate income rises by \$1 billion, aggregate expenditure on housing services rises by \$151 million. Is this a plausible figure?

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: HOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	334.6657	37.26625	8.980396	0.0000
DPI	0.150925	0.001665	90.65785	0.0000
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Durbin-Watson stat	0.337638	Prob(F-statistic)	0.000000

Possibly. It implies that 15 cents out of the marginal dollar are spent on housing. Housing is the largest category of consumer expenditure, so we would expect a substantial coefficient. Perhaps it is a little low.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: HOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	334.6657	37.26625	8.980396	0.0000
DPI	0.150925	0.001665	90.65785	0.0000
<b>PRELHOUS</b>	<b>-3.834387</b>	0.460490	-8.326764	0.0000

R-squared	0.996722	Mean dependent var	630.2830
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S.E. of regression	14.60740	Akaike info criteri	8.265274
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Log likelihood	-182.9687	F-statistic	6385.025
Durbin-Watson stat	0.337638	Prob(F-statistic)	0.000000

The coefficient of *PRELHOUS* indicates that a one-point increase in this price index causes expenditure on housing to fall by \$3.84 billion. It is not easy to determine whether this is plausible. At least the effect is negative.



# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: HOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	334.6657	37.26625	8.980396	0.0000
DPI	0.150925	0.001665	90.65785	0.0000
PRELHOUS	-3.834387	0.460490	-8.326764	0.0000

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Sum squared resid	8961.801	Schwarz criterion	8.385719
Log likelihood	-182.9687	F-statistic	6385.025
Durbin-Watson stat	0.337638	Prob(F-statistic)	0.000000

The constant has no meaningful interpretation. (Literally, it indicates that \$335 billion would be spent on housing services if aggregate income and the price series were both 0.)

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: HOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	334.6657	37.26625	8.980396	0.0000
DPI	0.150925	0.001665	90.65785	0.0000
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Log likelihood	-182.9687	F-statistic	6385.025
Durbin-Watson stat	0.337638	Prob(F-statistic)	0.000000

The explanatory power of the model appears to be excellent. The coefficient of *DPI* has a very high *t* statistic, that of price is also high, and  $R^2$  is close to a perfect fit.

$$HOUS = \beta_1 DPI^{\beta_2} PRELHOUS^{\beta_3} \nu$$

Constant elasticity functions are usually considered preferable to linear functions in models of consumer expenditure. Here  $\beta_2$  is the income elasticity and  $\beta_3$  is the price elasticity for expenditure on housing services.

$$HOUS = \beta_1 DPI^{\beta_2} PRELHOUS^{\beta_3} \nu$$

$$LGHOUS = \log \beta_1 + \beta_2 LGDPI + \beta_3 LGPRHOUS + \log \nu$$

We linearize the model by taking logarithms. We will regress *LGHOUS*, the logarithm of expenditure on housing services, on *LGDPI*, the logarithm of disposable personal income, and *LGPRHOUS*, the logarithm of the relative price index for housing services.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.005625	0.167903	0.033501	0.9734
<b>LGDP</b>	<b>1.031918</b>	0.006649	155.1976	0.0000
LGPRHOUS	-0.483421	0.041780	-11.57056	0.0000

R-squared	0.998583	Mean dependent var	6.359334
Adjusted R-squared	0.998515	S.D. dependent var	0.437527
S.E. of regression	0.016859	Akaike info criter	-5.263574
Sum squared resid	0.011937	Schwarz criterion	-5.143130
Log likelihood	121.4304	F-statistic	14797.05
Durbin-Watson stat	0.633113	Prob(F-statistic)	0.000000

Here is the regression output. The estimate of the income elasticity is 1.03. Is this plausible?

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.005625	0.167903	0.033501	0.9734
<b>LGDP</b>	<b>1.031918</b>	0.006649	155.1976	0.0000
LGPRHOUS	-0.483421	0.041780	-11.57056	0.0000

R-squared	0.998583	Mean dependent var	6.359334
Adjusted R-squared	0.998515	S.D. dependent var	0.437527
S.E. of regression	0.016859	Akaike info criter	-5.263574
Sum squared resid	0.011937	Schwarz criterion	-5.143130
Log likelihood	121.4304	F-statistic	14797.05
Durbin-Watson stat	0.633113	Prob(F-statistic)	0.000000

Probably. Housing is an essential category of consumer expenditure, and necessities generally have elasticities lower than 1. But it also has a luxury component, in that people tend to move to more desirable housing as income increases.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.005625	0.167903	0.033501	0.9734
LGDPI	1.031918	0.006649	155.1976	0.0000
<b>LGPRHOUS</b>	<b>-0.483421</b>	0.041780	-11.57056	0.0000
R-squared	0.998583	Mean dependent var	6.359334	
Adjusted R-squared	0.998515	S.D. dependent var	0.437527	
S.E. of regression	0.016859	Akaike info criter	-5.263574	
Sum squared resid	0.011937	Schwarz criterion	-5.143130	
Log likelihood	121.4304	F-statistic	14797.05	
Durbin-Watson stat	0.633113	Prob(F-statistic)	0.000000	

Thus an elasticity near 1 seems about right. The price elasticity is 0.48, suggesting that expenditure on this category is not very price elastic.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

=====  
Dependent Variable: LGHOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45  
=====

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.005625	0.167903	0.033501	0.9734
LGDP1	1.031918	0.006649	155.1976	0.0000
LGPRHOUS	-0.483421	0.041780	-11.57056	0.0000

=====

R-squared	0.998583	Mean dependent var	6.359334
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S.E. of regression	0.016859	Akaike info criter	-5.263574
Sum squared resid	0.011937	Schwarz criterion	-5.143130
Log likelihood	121.4304	F-statistic	14797.05
Durbin-Watson stat	0.633113	Prob(F-statistic)	0.000000

=====

Again, the constant has no meaningful interpretation.



# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample: 1959 2003

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.005625	0.167903	0.033501	0.9734
LGDPPI	1.031918	0.006649	155.1976	0.0000
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Sum squared resid	0.011937	Schwarz criterion	-5.143130
Log likelihood	121.4304	F-statistic	14797.05
Durbin-Watson stat	0.633113	Prob(F-statistic)	0.000000

The explanatory power of the model appears to be excellent.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Current and lagged values of the logarithm of disposable personal income

Year	<i>LGDP</i>	<i>LGDP</i> (-1)
1959	5.4914	—
1960	5.5426	5.4914
1961	5.5898	5.5426
1962	5.6449	5.5898
1963	5.6902	5.6449
1964	5.7371	5.6902
.....	.....	.....
.....	.....	.....
1999	6.8861	6.8553
2000	6.9142	6.8861
2001	6.9410	6.9142
2002	6.9679	6.9410
2003	6.9811	6.9679

Next, we will introduce some simple dynamics. Expenditure on housing is subject to inertia and responds slowly to changes in income and price. Accordingly we will consider specifications of the model where it depends on lagged values of income and price.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Current and lagged values of the logarithm of disposable personal income

Year	<i>LGDP</i>	<i>LGDP</i> (-1)
1959	5.4914	—
1960	5.5426	5.4914
1961	5.5898	5.5426
1962	5.6449	5.5898
1963	5.6902	5.6449
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.....	.....	.....
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1999	6.8861	6.8553
2000	6.9142	6.8861
2001	6.9410	6.9142
2002	6.9679	6.9410
2003	6.9811	6.9679

A variable  $X$  lagged one time period has values that are simply the previous values of  $X$ , and it is conventionally denoted  $X(-1)$ . Here  $LGDP(-1)$  has been derived from  $LGDP$ . You can see, for example, that the value of  $LGDP(-1)$  in 2003 is just the value of  $LGDP$  in 2002.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Current and lagged values of the logarithm of disposable personal income

Year	<i>LGDP</i>	<i>LGDP</i> (-1)
1959	5.4914	—
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.....	.....	.....
.....	.....	.....
1999	6.8861	6.8553
2000	6.9142	6.8861
2001	6.9410	6.9142
2002	6.9679	6.9410
2003	6.9811	6.9679

Similarly for the other years. Note that *LGDP*(-1) is not defined for 1959, given the data set. Of course, in this case, we could obtain it from the 1960 issues of the *Survey of Current Business*.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Current and lagged values of the logarithm of disposable personal income

Year	<i>LGDP</i>	<i>LGDP</i> (-1)	<i>LGDP</i> (-2)
1959	5.4914	—	—
1960	5.5426	5.4914	—
1961	5.5898	5.5426	5.4914
1962	5.6449	5.5898	5.5426
1963	5.6902	5.6449	5.5898
1964	5.7371	5.6902	5.6449
.....	.....	.....	.....
.....	.....	.....	.....
1999	6.8861	6.8553	6.8271
2000	6.9142	6.8861	6.8553
2001	6.9410	6.9142	6.8861
2002	6.9679	6.9410	6.9142
2003	6.9811	6.9679	6.9410

Similarly,  $LGDP(-2)$  is  $LGDP$  lagged 2 time periods.  $LGDP(-2)$  in 2003 is the value of  $LGDP$  in 2001, and so on. Generalizing,  $X(-s)$  is  $X$  lagged  $s$  time periods.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample (adjusted): 1960 2003

Included observations: 44 after adjusting endpoints

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.019172	0.148906	0.128753	0.8982
LGDP1 (-1)	1.006528	0.005631	178.7411	0.0000
LGPRHOUS (-1)	-0.432223	0.036461	-11.85433	0.0000

R-squared	0.998917	Mean dependent var	6.379059
Adjusted R-squared	0.998864	S.D. dependent var	0.421861
S.E. of regression	0.014218	Akaike info criter	-5.602852
Sum squared resid	0.008288	Schwarz criterion	-5.481203
Log likelihood	126.2628	F-statistic	18906.98
Durbin-Watson stat	0.919660	Prob(F-statistic)	0.000000

Here is a logarithmic regression of current expenditure on housing on lagged income and price. Note that EViews, in common with most regression applications, recognizes  $X(-1)$  as being the lagged value of  $X$  and there is no need to define it as a distinct variable.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample (adjusted): 1960 2003

Included observations: 44 after adjusting endpoints

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.019172	0.148906	0.128753	0.8982
LGDP1 (-1)	1.006528	0.005631	178.7411	0.0000
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R-squared	0.998917	Mean dependent var	6.379059
Adjusted R-squared	0.998864	S.D. dependent var	0.421861
S.E. of regression	0.014218	Akaike info criter	-5.602852
Sum squared resid	0.008288	Schwarz criterion	-5.481203
Log likelihood	126.2628	F-statistic	18906.98
Durbin-Watson stat	0.919660	Prob(F-statistic)	0.000000

The estimate of the lagged income and price elasticities are 1.01 and 0.43, respectively.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)
<i>LGDP</i>	1.03	—
	(0.01)	
<i>LGDP</i> (-1)	—	1.01
	(0.01)	
<i>LGDP</i> (-2)	—	—
<i>LGPRHOUS</i>	-0.48	—
	(0.04)	
<i>LGPRHOUS</i> (-1)	—	-0.43
	(0.04)	
<i>LGPRHOUS</i> (-2)	—	—
<i>R</i> <sup>2</sup>	0.9985	0.9989

The regression results will be summarized in a table for comparison. The results of the lagged-values regression are virtually identical to those of the current-values regression.



# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)	(3)
<i>LGDP</i>	1.03	—	—
(0.01)			
<i>LGDP</i> (-1)	—	1.01	—
(0.01)			
<i>LGDP</i> (-2)	—	—	0.98
(0.01)			
<i>LGPRHOUS</i>	-0.48	—	—
(0.04)			
<i>LGPRHOUS</i> (-1)	—	-0.43	—
(0.04)			
<i>LGPRHOUS</i> (-2)	—	—	-0.38
(0.04)			
<i>R</i> <sup>2</sup>	0.9985	0.9989	0.9988

So also are the results of regressing *LGHOUS* on *LGDP* and *LGPRHOUS* lagged two years.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)	(3)	(4)
<i>LGDP</i>	1.03—	—	0.33	
	(0.01)		(0.15)	
<i>LGDP</i> (-1)	—	1.01—	0.68	
	(0.01)	(0.15)		
<i>LGDP</i> (-2)	—	—	0.98—	
	(0.01)			
<i>LGPRHOUS</i>	-0.48	—	—	-0.09
	(0.04)	(0.17)		
<i>LGPRHOUS</i> (-1)	—	-0.43	—	-0.36
	(0.04)	(0.17)		
<i>LGPRHOUS</i> (-2)	—	—	-0.38	—
	(0.04)			
<i>R</i> <sup>2</sup>		0.9985	0.9989	0.9988
				<b>0.9990</b>

One approach to discriminating between the effects of current and lagged income and price is to include both in the equation. Since both may be important, failure to do so may give rise to omitted variable bias.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)	(3)	(4)
<i>LGDP</i>	1.03—	—	0.33	
	(0.01)		(0.15)	
<i>LGDP</i> (-1)	—	1.01—	0.68	
	(0.01)	(0.15)		
<i>LGDP</i> (-2)	—	—	0.98—	
	(0.01)			
<i>LGPRHOUS</i>	-0.48	—	—	-0.09
	(0.04)	(0.17)		
<i>LGPRHOUS</i> (-1)	—	-0.43	—	-0.36
	(0.04)	(0.17)		
<i>LGPRHOUS</i> (-2)	—	—	-0.38	—
	(0.04)			
<i>R</i> <sup>2</sup>		0.9985	0.9989	0.9988
				<b>0.9990</b>

With the current values of income and price, and their values lagged one year, we see that lagged income has a higher coefficient than current income. This is plausible, since we expect inertia in the response.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)	(3)	(4)
<i>LGDP</i>	1.03 (0.01)	—	0.33 (0.15)	
<i>LGDP</i> (-1)	—	1.01 (0.01)	0.68 (0.15)	
<i>LGDP</i> (-2)	—	—	0.98 (0.01)	
<i>LGPRHOUS</i>	-0.48 (0.04)	—	—	-0.09 (0.17)
<i>LGPRHOUS</i> (-1)	—	-0.43 (0.04)	—	-0.36 (0.17)
<i>LGPRHOUS</i> (-2)	—	—	-0.38 (0.04)	—
$R^2$	0.9985	0.9989	0.9988	0.9990

The price side of the model exhibits similar behavior.

## Alternative dynamic specifications, housing services

(4)

### Correlation Matrix

	LGDPPI	LGDPPI (-1)
LGDPPI	1.000000	0.999345
LGDPPI (-1)	0.999345	1.000000

<b><i>LGPRHOUS</i></b>	<b>-0.48</b>	<b>—</b>	<b>-0.09</b>
	<b>(0.04)</b>	<b>(0.17)</b>	
<b><i>LGPRHOUS(-1)</i></b>	<b>—</b>	<b>-0.43</b>	<b>-0.36</b>
	<b>(0.04)</b>	<b>(0.17)</b>	
<b><i>LGPRHOUS(-2)</i></b>	<b>—</b>	<b>—</b>	<b>-0.38</b>
	<b>(0.04)</b>		
<b><i>R</i><sup>2</sup></b>	<b>0.9985</b>	<b>0.9989</b>	<b>0.9988</b>

**0.9990**

However there is a problem of multicollinearity caused by the high correlation between current and lagged values. The correlation is particularly high for current and lagged income.

## Alternative dynamic specifications, housing services

(4)

### Correlation Matrix

```

=====
                LGPRHOUS  LGPRHOUS (-1)
=====
LGPRHOUS        1.000000    0.977305
LGPRHOUS (-1)   0.977305    1.000000
=====
    
```

<b><i>LGPRHOUS</i></b>	<b>-0.48</b>	<b>—</b>	<b>—</b>	<b>-0.09</b>
	<b>(0.04)</b>	<b>(0.17)</b>		
<b><i>LGPRHOUS(-1)</i></b>	<b>—</b>	<b>-0.43</b>	<b>—</b>	<b>-0.36</b>
	<b>(0.04)</b>	<b>(0.17)</b>		
<b><i>LGPRHOUS(-2)</i></b>	<b>—</b>	<b>—</b>	<b>-0.38</b>	<b>—</b>
	<b>(0.04)</b>			
<b><i>R</i><sup>2</sup></b>	<b>0.9985</b>	<b>0.9989</b>	<b>0.9988</b>	<b>0.9990</b>

The correlation is also high for current and lagged price.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)	(3)	(4)
<i>LGDP</i>	1.03 (0.01)	—	0.33 (0.15)	
<i>LGDP</i> (-1)	—	1.01 (0.01)	0.68 (0.15)	
<i>LGDP</i> (-2)	—	—	0.98 (0.01)	
<i>LGPRHOUS</i>	-0.48 (0.04)	—	—	-0.09 (0.17)
<i>LGPRHOUS</i> (-1)	—	-0.43 (0.04)	—	-0.36 (0.17)
<i>LGPRHOUS</i> (-2)	—	—	-0.38 (0.04)	—
$R^2$		0.9985	0.9989	0.9988
				<b>0.9990</b>

Notice how the standard errors have increased. The fact that the coefficients seem plausible is probably just an accident.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Alternative dynamic specifications, housing services

<i>Variable</i>	(1)	(2)	(3)	(4)	(5)
<i>LGDP</i>	1.03 (0.01)	—	0.33 (0.15)	0.29 (0.14)	
<i>LGDP</i> (-1)	—	1.01 (0.01)	0.68 (0.15)	0.22 (0.20)	
<i>LGDP</i> (-2)	—	—	0.98 (0.01)	0.49 (0.13)	
<i>LGPRHOUS</i>	-0.48 (0.04)	—	—	-0.09 (0.17)	-0.28
<i>LGPRHOUS</i> (-1)	—	-0.43 (0.04)	—	-0.36 (0.30)	0.23
<i>LGPRHOUS</i> (-2)	—	—	-0.38 (0.04)	—	-0.38
<i>R</i> <sup>2</sup>		0.9985	0.9989	0.9988	0.9990
					<b>0.9993</b>

If we add income and price lagged two years, the results become even more erratic. For a category of expenditure such as housing, where one might expect long lags, this is clearly not a constructive approach to determining the lag structure.



# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

Despite the problem of multicollinearity, we may be able to obtain relatively precise estimates of the long-run elasticities with respect to income and price.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

The usual way of investigating the long-run relationship between  $Y$  and  $X$  is to perform an exercise in comparative statics. One first determines how equilibrium  $\check{Y}$  would be related to equilibrium  $\check{X}$ , if the process ever reached equilibrium.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
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$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

One then evaluates the effect of a change in equilibrium  $\check{X}$  on equilibrium  $\check{Y}$ .

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

In the model with two lags shown,  $(\beta_2 + \beta_3 + \beta_4)$  is a measure of the long-run effect of  $X$ . We contrast this with the short-run effect, which is simply  $\beta_2$ , the impact of current  $X_t$  on  $Y_t$ .

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

We can calculate the long-run effect from the point estimates of  $\beta_2$ ,  $\beta_3$ , and  $\beta_4$  in the original specification. The estimate of the sum may be quite stable, even though the estimates of the individual coefficients may be subject to multicollinearity.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

The table presents an example of this. It gives the sum of the income and price elasticities for the five specifications of the logarithmic housing demand function considered earlier. The estimates of the long-run elasticities are very similar.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

$$Y_t = \beta_1 + (\beta_2 + \beta_3 + \beta_4) X_t - \beta_3 (X_t - X_{t-1}) - \beta_4 (X_t - X_{t-2}) + u_t$$

If we are estimating long-run effects, we need standard errors as well as point estimates. The most straightforward way of obtaining the standard error is to reparameterize the model. In the case of the present model, we could rewrite it as shown.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

$$Y_t = \beta_1 + (\beta_2 + \beta_3 + \beta_4) X_t - \beta_3 (X_t - X_{t-1}) - \beta_4 (X_t - X_{t-2}) + u_t$$

The point estimate of the coefficient of  $X_t$  will be the sum of the point estimates of  $\beta_2$ ,  $\beta_3$ , and  $\beta_4$  in the original specification and so the standard error of that coefficient is the standard error of the estimate of the long-run effect.



# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

## Estimates of long-run income and price elasticities

Specification	(1)	(2)	(3)	(4)	(5)
Sum of income elasticities	1.03	1.01	0.98	1.01	1.00
Sum of price elasticities	-0.48	-0.43	-0.38	-0.45	-0.43

$$Y_t = \beta_1 + \beta_2 X_t + \beta_3 X_{t-1} + \beta_4 X_{t-2} + u_t$$

$$\check{Y} = \beta_1 + \beta_2 \check{X} + \beta_3 \check{X} + \beta_4 \check{X} = \beta_1 + (\beta_2 + \beta_3 + \beta_4) \check{X}$$

$$Y_t = \beta_1 + (\beta_2 + \beta_3 + \beta_4) X_t - \beta_3 (X_t - X_{t-1}) - \beta_4 (X_t - X_{t-2}) + u_t$$

Since  $X_t$  may well not be highly correlated with  $(X_t - X_{t-1})$  or  $(X_t - X_{t-2})$ , there may not be a problem of multicollinearity and the standard error may be relatively small.

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample (adjusted): 1961 2003

Included observations: 43 after adjusting endpoints

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.046768	0.133685	0.349839	0.7285
LGDP1	1.000341	0.006997	142.9579	0.0000
X1	-0.221466	0.196109	-1.129302	0.2662
X2	-0.491028	0.134374	-3.654181	0.0008
LGPRHOUS	-0.425357	0.033583	-12.66570	0.0000
P1	-0.233308	0.298365	-0.781955	0.4394
P2	0.378626	0.175710	2.154833	0.0379

R-squared 0.999265 Mean dependent var 6.398513  
 Adjusted R-squared 0.999143 S.D. dependent var 0.406394  
 Total sum of squares 5.076007

$$Y_t = \beta_1 + (\beta_2 + \beta_3 + \beta_4)X_t - \beta_3(X_t - X_{t-1}) - \beta_4(X_t - X_{t-2}) + u_t$$

The output shows the result of fitting the reparameterized model for housing with two lags (Specification (5) in the table).  $X1 = LGDPI - LGDPI(-1)$ ,  $X2 = LGDPI - LGDPI(-2)$ ,  $P1 = LGPRHOUS - LGPRHOUS(-1)$ , and  $P2 = LGPRHOUS - LGPRHOUS(-2)$ .

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample (adjusted): 1961 2003

Included observations: 43 after adjusting endpoints

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.046768	0.133685	0.349839	0.7285
LGDP1	1.000341	0.006997	142.9579	0.0000
X1	-0.221466	0.196109	-1.129302	0.2662
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P1	-0.233308	0.298365	-0.781955	0.4394
P2	0.378626	0.175710	2.154833	0.0379

R-squared 0.999265 Mean dependent var 6.398513  
 Adjusted R-squared 0.999143 S.D. dependent var 0.406394  
 Sum of squared residuals 0.011000 Akaike info criterion 5.076007

$$Y_t = \beta_1 + (\beta_2 + \beta_3 + \beta_4)X_t - \beta_3(X_t - X_{t-1}) - \beta_4(X_t - X_{t-2}) + u_t$$

As expected, the point estimates of the coefficients of *LGDP1* and *LGPRHOUS*, 1.00 and -0.43, are the sum of the point estimates of the coefficients of the current and lagged terms in Specification (5).

# TIME SERIES MODELS: STATIC MODELS AND MODELS WITH LAGS

Dependent Variable: LGHOUS

Method: Least Squares

Sample (adjusted): 1961 2003

Included observations: 43 after adjusting endpoints

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.046768	0.133685	0.349839	0.7285
LGDP1	1.000341	0.006997	142.9579	0.0000
X1	-0.221466	0.196109	-1.129302	0.2662
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P1	-0.233308	0.298365	-0.781955	0.4394
P2	0.378626	0.175710	2.154833	0.0379

R-squared 0.999265 Mean dependent var 6.398513  
 Adjusted R-squared 0.999143 S.D. dependent var 0.406394  
 Total Sum of Squares 2.011000 Residual Sum of Squares 0.076007

$$Y_t = \beta_1 + (\beta_2 + \beta_3 + \beta_4)X_t - \beta_3(X_t - X_{t-1}) - \beta_4(X_t - X_{t-2}) + u_t$$

Also as expected, the standard errors, 0.01 and 0.03, are much lower than those of the individual coefficients in Specification (5).

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